

INCOHERENCE CORRECTION STRATEGIES IN STATISTICAL MATCHING

Andrea Capotorti and Barbara Vantaggi

26-VII-2011

ABOUT OURSELVES...

Both of us come from the so called ... “0-probability Italian school” !!!

and in particular:

- Barbara is Associate Professor at Roma “la Sapienza”;
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Probability doesn't exist ! (de Finetti's provocative statement)

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Probabilistic evaluations are commonly conditional and partial
(our “mantra”)

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Common probabilistic models are imprecise ! (SIPTA “keystone”)

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The present paper comes from the merging (hence it's (*meta*)**matching** ! ;-)) of two our previous ISIPTA's contributions:

- B. Vantaggi. The role of coherence for the integration of different sources. *ISIPTA'05* (Pittsburgh, USA):
- A. Capotorti, G. Regoli and F. Vattari. On the use of a new discrepancy measure to correct incoherent assessments and to aggregate conflicting opinions based on imprecise conditional probabilities. *ISIPTA'09* - Durham (UK).

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IN PARTICULAR...

- we deal with the managing of inconsistencies inside the Statistical Matching (**integration of sources**) framework;
 - when logical relations among the variables are present incoherence can arise in the probability evaluations
- different methods can be used to remove such incoherences:
 - maximize the “partial likelihood function” on the base of observed data;
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COHERENT EXTENSION

To adjust the initially incoherent assessment $(\mathcal{E}, \mathbf{p})$ it is possible to determine a coherent sub-assessment $(\mathcal{G}, \mathbf{p}_{|\mathcal{G}})$ with maximal cardinality and coherently extend it to the rest $\mathcal{F} = \mathcal{E} \setminus \mathcal{G}$ by the generalized Bayesian updating scheme obtaining an imprecise sub-assessment

$$(\mathcal{F}, [\underline{\mathbf{p}}_{\mathcal{F}}, \overline{\mathbf{p}}_{\mathcal{F}}]).$$

Note that **inference on decision targets** can be performed again through the generalized Bayesian updating scheme but applied to imprecise evaluations.

Whenever too vague, inference bounds can be eventually reduced to coherent cores, i.e. *total* coherent subintervals with highest degree of support.

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(PSEUDO)DISTANCES AMONG PROBABILITY DISTRIBUTIONS

Given two conditional assessments $\mathbf{p} = [p_1, \dots, p_n]$ and $\mathbf{q} = [q_1, \dots, q_n]$ on the same set of conditional events $\mathcal{E} = [E_1|H_1, \dots, E_n|H_n]$, the most widely adopted divergencies among them are:

$$\textcircled{1} \quad L1(\mathbf{p}, \mathbf{q}) = \sum_{i=1}^n |q_i - p_i|;$$

$$\textcircled{2} \quad L2(\mathbf{p}, \mathbf{q}) = \sum_{i=1}^n (q_i - p_i)^2;$$

$$\textcircled{3} \quad KL(\mathbf{p}, \mathbf{q}) = \sum_{i=1}^n (q_i \ln(q_i/p_i) - q_i + p_i).$$

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OUR MAIN TOOL

... but for partial conditional probability assessments $\mathbf{p} \in (0, 1)^n$ on \mathcal{E} recently we tailored the following “discrepancy”

$$\Delta(\mathbf{p}, \alpha) = \sum_{i|\alpha(H_i) > 0} \alpha(H_i) \left(q_i \ln \frac{q_i}{p_i} + (1 - q_i) \ln \frac{(1 - q_i)}{(1 - p_i)} \right),$$

where \mathbf{q}_α is an assessment on \mathcal{E} induced by the probability mass distribution α ,

that now we specialize for Statistical Matching with a mixture of discrepancies $\Delta_{mix}(\mathbf{p}, \{\alpha_j\}_j) \dots$

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INTEGRATION OF SOURCES IN A COHERENT SETTING

$(X_1, Y_1), \dots, (X_{n_A}, Y_{n_A})$ and $(X_{n_A+1}, Z_{n_A+1}), \dots, (X_{n_A+n_B}, Z_{n_A+n_B})$
 two random samples (with a finite range) related to two sources A
 and B concerning the same population of interest, and drawn
 according to the same sampling scheme, with

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 exchangeable, as well as the sequence

$X_1, \dots, X_{n_A}, X_{n_A+1}, \dots, X_{n_A+n_B}$

We can elicit from the two files the relevant probability values:

$$y_{j|i} = P_{Y|(X=x_i)}(Y = y_j) \quad z_{k|i} = P_{Z|(X=x_i)}(Z = z_k) \quad x_i = P_X(X = x_i)$$

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... AND THE MAIN PROBLEM

Given $\mathcal{Y}_{j|i}, \mathcal{Z}_{k|i}, \mathcal{X}_i$, for any i, j, k , with $(\mathcal{Y}_{j|i}, \mathcal{X}_i)$ and $(\mathcal{Z}_{k|i}, \mathcal{X}_i)$ **separately coherent** and with some logical constraints among the variables Y and Z , **incoherence** in the **whole assessment**

$$(\mathcal{E}, \mathbf{p}) \text{ with } \mathcal{E} = \left\{ \begin{array}{l} (X = x_i), (Y = y_j) | (X = x_i), (Z = z_k) | (X = x_i) \\ \text{for any } x_i, y_j, z_k \end{array} \right\},$$

$$\mathbf{p} = \{x_i, \mathcal{Y}_{j|i}, \mathcal{Z}_{k|i}\}_{i,j,k} .$$

can *localize* only in association to elements of \mathcal{E} **with the same conditioning event** $(X = x_i)$.

A MIXTURE OF DISCREPANCIES

Hence our discrepancy $\Delta(\mathbf{p}, \alpha)$ can be reformulated into

$$\begin{aligned} \Delta_{mix}(\mathbf{p}, \{\alpha_i\}_i) &= \sum_i x_i \left[\sum_j \left(q_{j|i}^{\alpha_i} \ln \frac{q_{j|i}^{\alpha_i}}{y_{j|i}} + (1 - q_{j|i}^{\alpha_i}) \ln \frac{(1 - q_{j|i}^{\alpha_i})}{(1 - y_{j|i})} \right) + \right. \\ &\quad \left. + \sum_k \left(q_{k|i}^{\alpha_i} \ln \frac{q_{k|i}^{\alpha_i}}{z_{k|i}} + (1 - q_{k|i}^{\alpha_i}) \ln \frac{(1 - q_{k|i}^{\alpha_i})}{(1 - z_{k|i})} \right) \right] \end{aligned}$$

where each distribution α_i works just on the sample space spanned by the conditional events $\{(Y = y_j)|(X = x_i), (Z = z_k)|(X = x_i)\}$, it is constrained to fulfill the normalizing condition $\alpha_i(X = x_i) = x_i$, and generates the conditional probabilities

$$q_{j|i}^{\alpha_i} = \frac{\alpha_i(Y = y_j)}{\alpha_i(X = x_i)} \quad q_{k|i}^{\alpha_i} = \frac{\alpha_i(Z = z_k)}{\alpha_i(X = x_i)}.$$

OUR RESULTS

We applied the three methodologies (**likelihood maximization**, **coherent extension and distances minimizations**) to real data representing a subset of employees extracted from a pilot survey of the Italian Population and Household Census that induce an **incoherent** initial assessment.

Three categorical variables have been analyzed: Age, Educational Level and Professional Status.

And we obtained...

	$S_1 A_4$	$S_2 A_4$	$S_3 A_4$	$E_1 A_4$	$E_2 A_4$	$E_3 A_4$	$E_4 A_4$	$S_3 E_4$
p	0.6667	0.1111	0.2222	0.6667	0	0.2000	0.1333	\emptyset
$L1 _{\mathcal{F}}$	0.2222	-	0.6667	0.6667	-	-	-	[0,0.6285]
$L1 _{A_4}$	0.5266	0	0.4734	0.4734	0	0.2836	0.2431	[0,0.6234]
$L2 _{A_4}$	0.5333	0.0389	0.4278	0.4278	0.0389	0.3	0.2333	[0,0.6238]
$KL _{A_4}$	0.4856	0.1179	0.3965	0.3965	0.1179	0.2914	0.1942	[0,0.6257]
Δ_{mix}	0.4985	0.0939	0.4077	0.4077	0.0939	0.2943	0.2042	[0,0.6252]
ML	0.4286	0.0714	0.5000	0.5000	0	0.3000	0.2000	[0,0.6254]
$IP_{\mathcal{E} \setminus \mathcal{F}}$ core	[0, 0.2222]	-	[0.6667 0.8889]	-	-	-	-	[0,0.6386] [0.0017,0.6286]
$IP_{\mathcal{E} \setminus \{\cdot A_4\}}$ core	[0, 1]	[0, 1]	[0, 1]	[0, 1]	[0, 1]	[0, 1]	[0, 1]	[0,0.6607] [0,0.6349]

... details at the poster ! but here we can say...

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CONCLUSION

- conditional events involved in incoherencies are those conditioned to the age status A_4 , in particular, the minimal set is $\mathcal{F} = \{E_1|A_4, S_1|A_4, S_3|A_4\}$;
- $L1|_{\mathcal{F}}$ and $IP_{\mathcal{E} \setminus \mathcal{F}}$ perform quite well: even though a drastic change on the probability values, they induce quite reasonable inference bounds;
- $L1|_{A_4}$ and ML give similar results and in particular they leave to 0 the probability of $E_2|A_4$ since the absence of observations in the original data;
- others “precise” adjustments have all quite similar behaviors;
- Δ_{mix} has the advantage of automatically localize of the scenarios where the adjustment can be performed;
- the wider imprecise correction $IP_{\mathcal{E} \setminus \{\cdot|A_4\}}$, being the one with less assumption requirement, surely performs worst.

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- Δ_{mix} has the advantage of automatically localize of the scenarios where the adjustment can be performed;
- the wider imprecise correction $IP_{\mathcal{E}\setminus\{ \cdot | A_4 \}}$, being the one with less assumption requirement, surely performs worst.

CONCLUSION

- conditional events involved in incoherencies are those conditioned to the age status A_4 , in particular, the minimal set is $\mathcal{F} = \{E_1|A_4, S_1|A_4, S_3|A_4\}$;
- $L1|_{\mathcal{F}}$ and $IP_{\mathcal{E}\setminus\mathcal{F}}$ perform quite well: even though a drastic change on the probability values, they induce quite reasonable inference bounds;
- $L1|_{A_4}$ and ML give similar results and in particular they leave to 0 the probability of $E_2|A_4$ since the absence of observations in the original data;
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A FINAL REMARK...

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... but till now mainly influenced by government financial and motivational destruction of public and free universities



ROME-PERUGIA AXIS...

